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## Particulate air pollutants and trajectories of depressive symptoms in older women.

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### Abstract

**Objectives:** Although several environmental factors contribute to the etiology of late-life depressive symptoms, the role of ambient air pollution has been understudied. Experimental data support the neurotoxicity of airborne particulate matter with aerodynamic diameter of 2.5  $\mu\text{m}$  (PM<sub>2.5</sub>), but it remains unclear whether long-term exposure is associated with late-life depressive symptoms. Our secondary aim was to explore whether the observed associations between exposure and depressive symptoms are explained by dementia risk.

**Design, Setting, and Participants:** Prospective community-dwelling cohort study from the Women's Health Initiative Study of Cognitive Aging (1999–2010). Our analyses included 1,989 older women (baseline age 73.3+3.75) with no prior depression or cognitive impairment.

**Measurements:** Participants completed annual assessments of depressive symptoms (15-item Geriatric Depression Scale). Average ambient PM<sub>2.5</sub> exposure at the residential location was estimated by spatiotemporal modeling for the three-years preceding each neuropsychological

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**Previous Presentation:** Portions of this work have been presented previously at the Joint Annual Meeting of the International Society of Exposure Science and the International Society of Environmental Epidemiology taking place August 26 – August 30, 2018, in Ottawa CA.

assessment. Participants underwent separate annual examinations for incident dementia defined by DSM-IV. Latent-class mixture models examined the association between PM<sub>2.5</sub> and identified trajectories of symptoms.

**Results:** Six trajectories of depressive symptoms were identified. Across all women, PM<sub>2.5</sub> exposure was positively associated with depressive symptoms. The effect was especially strong in two clusters with sustained depressive symptoms (n=625 sustained-mild [31%]; n=125 sustained-moderate; [6%]). Among those with sustained-moderate symptoms, the estimated adverse effect of PM<sub>2.5</sub> exposure was greater than that of hypertension. Among women without dementia, associations were modestly attenuated.

**Conclusions:** Long-term exposure to ambient fine particles was associated with increased depressive symptoms among older women without prior depression or cognitive impairment.

### Keywords

depressive symptoms; air pollution; epidemiology

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### Objective

Late-life depressive symptoms are common, disproportionately affect women (1), impair functioning and quality of life, and are a significant burden to public health (2), even when the total number of symptoms does not meet criteria for a clinical diagnosis (3). Etiologies of depressive symptoms that emerge in later life (after age 65) likely differ from the causes of mood symptoms with an earlier-life onset (4). Both clinical and epidemiologic data has shown that late-life depression is heterogenous in nature with multiple contributing factors including cerebrovascular disease (5), neurodegeneration, and medical comorbidities (6). While twin studies suggest that late-life depressive symptoms are heritable, environmental factors explain the majority of the variance (7). However, there has been limited longitudinal environmental epidemiologic research on depressive symptoms that considers the distinct trajectories of symptoms over time (8). Previous research largely focused on psychosocial factors and depressive symptoms (9), and the potential role of physical environments has been understudied.

There is increasing evidence that ambient air pollutants, especially fine particulate matter (PM<sub>2.5</sub>; aerodynamic diameter <2.5 μm), is associated with accelerated brain aging (10). For instance, late-life exposure to PM<sub>2.5</sub> was associated with increased risk for dementia among older women in the US (11). Experimental models also demonstrate airborne particle exposure can accelerate the neurodegenerative processes (11–13). Although neurotoxicological studies have shown depressive-like behaviors in mice exposed to particles (14–16), convincing epidemiologic data are lacking. Few longitudinal studies have examined the long-term associations between PM<sub>2.5</sub>, depressive symptomatology and onset of depression, but findings were mixed (17–20). These studies suffered from major methodological weaknesses, including short follow-up time with only two repeated assessments (18,19), ignoring the heterogeneity in trajectories of depressive symptoms over time (18–20), utilizing medical records for the ascertainment of depression (17,20), and inadequate control of potential confounders (20). Past research has not excluded individuals

with a past history of depression so results may be biased due to potential confounding by other environmental exposure that have affected emotional health before the study baseline (19). Lastly, no published data has offered insights into the possible mechanistic pathways underlying the hypothesized associations.

Here, we conducted a longitudinal study to examine the association between air pollution and late-life depressive symptoms assessed annually (1999–2010) in a geographically-diverse sample of community-dwelling older women without prior depression or cognitive impairment. Our primary aim was to investigate whether long-term exposure to  $PM_{2.5}$  in late-life affected the trajectories of depressive symptoms, while rigorously controlling for potential confounders. We hypothesized that exposure to  $PM_{2.5}$  will be positively associated with depressive symptoms in later life. Our secondary aim was to explore whether any associations between exposure and depressive symptoms may be explained by dementia risk, regional brain volumes, or magnetic resonance imaging (MRI) measures of cerebrovascular damage that are known to be associated with  $PM_{2.5}$  exposure and/or depressive symptoms in late life. We hypothesized that dementia risk, regional brain volumes, and MRI measures of cerebrovascular damage would partially explain any putative associations between  $PM_{2.5}$  exposure and late life depressive symptoms.

## Methods

### Study Population

This study included 1,989 community-dwelling older women (aged  $73.3 \pm 3.8$  years old at baseline) without prior depression or dementia who were enrolled in the Women's Health Initiative Study of Cognitive Aging (WHISCA;  $n=2,304$ ) (21), an ancillary study to the Women's Health Initiative Memory Study (WHIMS) (22) which was an ancillary study to the Women's Health Initiative Hormone Clinical Trial (23). These women completed annual (1999–2010) neuropsychological assessments, beginning on average three years after randomization in the Women's Health Initiative hormone therapy trial. Excluded from the present study (see Figure 1) were women who reported a history of depression ( $n=221$ ) before the initial WHISCA assessment and those with missing data on covariates of interest ( $n=92$ ). Our analyses also included a subsample of WHISCA participants ( $n=947$ ) who underwent a single brain MRI scan in the Women's Health Initiative Memory Study-MRI (WHIMS-MRI) study between 2005–2006 (24). A detailed description of these original study cohorts was given in the Supplemental Methods. This study was approved by the institutional review boards at all institutions and all participants provided written informed consent.

### Assessment of depressive symptoms

Depressive symptoms were assessed at WHISCA baseline and at each annual follow-up (up to 9 assessments) using the 15-item Geriatric Depression Scale (GDS-15) (25,26). WHISCA participants were coded as having a past history of depression if they stated yes to both of the depression questions from the National Institute of Mental Health Diagnostic Interview Schedule at the WHIMS baseline that took place between 1996–1999 (27).

### Assessment of ambient PM<sub>2.5</sub>

Detailed procedures for PM<sub>2.5</sub> estimation have previously been reported (11, 28). Briefly, residential addresses were collected at WHI enrollment and at each annual assessment. Standardized protocols (29) were followed for geocoding residential addresses. Daily ambient concentration of PM<sub>2.5</sub> at each residential location was estimated using the Bayesian Maximum Entropy-based spatiotemporal modeling method. This approach integrates nationwide monitoring data from the U.S. Environmental Protection Agency Air Quality System and the output of chemical transport models. Estimates of daily PM<sub>2.5</sub> exposures were statistically cross-validated (average Pearson's  $R^2 = .70$ ). The resulting exposure estimates were then aggregated in creating time-dependent estimates of average PM<sub>2.5</sub> exposure for the 3-years preceding each respective WHISCA assessment.

### Classification of dementia

Incident cases of all-cause dementia were defined using Diagnostic and Statistical Manual of Mental Disorders, Fourth edition (DSM-4) (30) criteria. Classification of dementia was based on annual screening of cognitive function, neuropsychological and functional assessment, and collection of clinical data to rule out possible reversible causes of cognitive declines through continued WHIMS participation. Annual screenings of all WHIMS participants were completed through administration of the Modified Mini Mental Status (3MS)(31) examination. Beginning in 2008, and continuing to the present, annual cognitive assessment were conducted by telephone using a validated battery (32,33) via participation in the WHIMS-Epidemiology of Cognitive Health Outcomes study.

### Structural MRI assessment

Participants underwent one structural MRI, collected by standardized protocols (24). Measures of brain volume (in cm<sup>3</sup>) included: total volume of normal-appearing white matter, prefrontal grey matter volume, hippocampal volume and total volumes of small-vessel-ischemic-diseases. These variables were selected because previous studies suggest they were brain areas affected by PM<sub>2.5</sub> exposure (34,35) or associated with late-life depression (5,36).

### Assessment of covariates

A structured questionnaire was administered at the WHIMS baseline (between 1996–1999) to gather information on the time-independent covariates of demographics (age, race/ethnicity), geographic region of residence (Northeast, South, Midwest, and West), socioeconomic status (education; family income; employment status), lifestyle factors (smoking; alcohol use; physical activities), clinical characteristics (postmenopausal hormone treatment, history of CVD (including previous coronary heart, stroke, or transient ischemic attack), hypertension (defined as elevated blood pressure or use of antihypertensive medication), and diabetes mellitus (defined as physician diagnosis plus oral medications, or insulin therapy). Good reliability and validity of the self-reported medical histories and the physical measures have been previously documented (37). Estimates were adjusted for global cognitive ability by including the 3MS (38) as a covariate in all analyses. The 3MS is a reliable and valid measure of global cognitive ability and was administered at the WHIMS baseline assessment.

## Statistical analysis

Latent class mixed models (LCMM) (39) with a latent process transformation applied to the continuous score of the GDS-15 (to account for non-Gaussian nature of depressive symptoms; see Supplemental Figure 1 for graph of the transformation function) were fit to identify groups of women who had similar trajectories of depressive symptoms over time (See a more detailed description in the Supplemental Methods). A higher score represented more depressive symptoms and the function linking continuous GDS-15 scores to the transformed score is provided in Supplemental figure 1. Years since WHISCA baseline was utilized as the timescale. We started with a base LCMM that included a quadratic function of time (years since WHISCA baseline squared) and attrition (time-dependent covariate if the participant dropped out at the subsequent visit; 0=no, 1=yes) as fixed effects to identify the number of classes. We first fit a one-class model and sequentially increased the number of estimated latent classes. The decision of the total number of latent classes to be retained was made by a combination of overall model fit as evidenced by the Bayesian Information Criterion (BIC), the interpretability of the latent classes, the clinical meaningfulness of the class, and how well the latent classes differentiated between participants. See the Supplemental methods for a more detailed description of the approach to decide the number of significant latent classes to retain.

A second series of LCMM analyses were fit to examine the associations between PM<sub>2.5</sub> exposure and depressive symptoms. While extracting the same number of latent classes as identified from the base LCMM, we first examined whether PM<sub>2.5</sub> exposure had a global effect on the transformed GDS-15 scores, assuming a common exposure effect across all identified trajectories after adjusting for demographic features, socioeconomic status, lifestyle factors, and clinical characteristics. In the subsequent LCMM adjusting for the same set of covariates, the putative association between PM<sub>2.5</sub> and depressive symptoms was then re-examined as a class-specific effect, which allowed the association with PM<sub>2.5</sub> to vary in magnitude by latent class.

## Secondary analyses

To explore whether dementia risk statistically explained the observed PM<sub>2.5</sub> exposure effects, the first set of secondary analyses were conducted only using the data collected in the subsample (n=1,811) who had not developed dementia by December 2015, as determined by standardized protocols and central adjudication (11). Participants were assigned *a priori* to latent class membership based on the same class membership already determined in the full-sample LCMM with PM<sub>2.5</sub> included. We examined the global and class-specific PM<sub>2.5</sub> associations and compared the magnitude of these effect estimates to the results from the full model.

Exploratory analyses were further conducted in the WHISCA participants with brain MRI data (n=947), to examine whether MRI estimates of structural brain volumes statistically explained the observed association between PM<sub>2.5</sub> and depressive symptoms. We examined the global PM<sub>2.5</sub> effect in this subsample, as well as the class-specific exposure effect, which was only tested among women belonging to classes with significant associations as

identified from the primary analyses. Subsequent LCMMs were constructed separately by adding respective MRI variables into the model.

### Sensitivity analyses

Lastly, two additional sets of sensitivity analyses were conducted: 1) using an alternative approach to defining prior depression; and 2) further adjusting for neighborhood socioeconomic characteristics (40) (detailed in the Supplemental methods). All models were fit with the LCMM program (39) in R version 3.5.0 (41).

## Results

### Participant characteristics, PM<sub>2.5</sub> exposure distribution and trajectories of depressive symptoms

The majority of these older women (N=1,989; aged 73.3±3.8 years) were Caucasian (91%) and had more than a high-school education (74%). Table 1 compares the distribution of 3-year average PM<sub>2.5</sub> exposure estimates by population characteristics. Participants exposed to higher levels of PM<sub>2.5</sub> were more likely to self-identify as African-American or Hispanic, reside in the Midwest or West, report higher household incomes ( > \$75,000) or prior histories of cardiovascular disease.

Base LCMM identified six trajectories of depressive symptoms (Supplemental Table 1), including: 779 with minimum GDS-15 scores throughout the study period (39% minimal), 228 with depressive symptoms emerging early in follow-up (11% early-emerging), 159 with depressive symptoms increasing during the latter part of the follow-up period (8% late-emerging), 644 with mildly elevated GDS-15 scores sustained during the follow-up (32% sustained-mild), 70 starting with elevated GDS-15 scores that decreased over the follow-up (4% decreasing), and 109 with sustained-moderate GDS-15 scores (5% sustained-moderate). Figure 2 (panel A) depicts the estimated mean GDS-15 scores over time with 95% confidence intervals for each latent class. A reference line was included on this graph at a score of 5 or higher to aid in clinical interpretation of these classes as this score is a commonly used cutoff for clinically significant depressive symptoms. Supplemental Figure 2 presents the estimated mean GDS-15 score with individual scores over time. The odds of correct classification were over 5 for all six latent classes, and, with the exception of the late-emerging class (PP = .65), had average posterior probabilities over .70. Table 2 presents the study variables by the six identified latent classes. Clinically meaningful differences between classes were present as women with sustained-mild or sustained-moderate symptoms tended to be in worse physical health with higher proportions of hypertension, cardiovascular disease, and diabetes at baseline, less frequently engaged in moderate or strenuous physical activity, and were more likely to drink alcohol compared to women in other trajectories.

### Associations between PM<sub>2.5</sub> exposure and GDS trajectories

Results from the multivariate-adjusted LCMM analyses are presented in Table 3. In the LCMM including PM<sub>2.5</sub> with adjustment for multiple covariates, a six-class solution still provided the best fit to the data. In models estimating the global PM<sub>2.5</sub> effect constrained

across the six identified classes, we found significantly more depressive symptoms ( $\beta=.043$ ;  $p=.039$ ) among older women living in locations with elevated  $PM_{2.5}$  (~ by  $3.45 \mu\text{g}/\text{m}^3$ ). When examining class-specific effects (all  $PM_{2.5}$  estimates presented in the left panel of Table 3), higher  $PM_{2.5}$  exposure was positively associated with depressive symptoms only in women who had sustained-mild ( $\beta=.096$ ;  $p=.023$ ) and sustained-moderate symptoms ( $\beta=.171$ ;  $p=.024$ ), but the strengths of associations within these two classes were 2–4 times greater than the estimated global main effect. Figure 2 (panel B) provides the estimated average GDS-15 score among women in the trajectories of sustained-mild or -moderate depressive symptoms, separately by high (upper quartile), median, and low (lower quartile)  $PM_{2.5}$  levels as compared with the corresponding effect estimates of selected covariates (estimated at the mean level of  $PM_{2.5}$ ). Estimated depressive symptoms for women exposed to the median  $PM_{2.5}$  and for those who had diabetes and hypertension at baseline were also plotted to put the  $PM_{2.5}$  effect into the clinical context for comparison. Among those with sustained-moderate symptoms, the magnitude of the adverse  $PM_{2.5}$  effect estimate (per 1-interquartile exposure range) was greater than that of hypertension.

### Secondary analyses

Long-term  $PM_{2.5}$  exposure was associated with increased depressive symptoms in the dementia-free sample ( $n = 1,811$ ). Compared to the corresponding LCMM results in the full sample, the estimates of class-specific  $PM_{2.5}$  effects were similar among WHISCA participants without dementia (right panel of Table 3), although the strength of association was slightly diminished and more precise in the subgroup with sustained-moderate depressive symptoms. The global effect estimate was attenuated, but remained statistically significant ( $\beta=.038$ ;  $p = .006$ ). Interestingly, the class-specific LCMM also revealed the statistically significant association with  $PM_{2.5}$  ( $\beta=.175$ ;  $p < .001$ ) among WHISCA participants who had late-emerging depressive symptoms but did not have dementia.

### Exploratory analyses

Among the subsample of WHISCA participants with brain MRI data ( $n=974$ ), both the global and class-specific  $PM_{2.5}$  exposure effects were statistically significant (Supplemental Table 2; Model-1). These associations were largely unchanged and remained statistically significant after we further accounted for various MRI-measured regional brain volumes (Supplemental Table 2; Model-2 to Model-5).

### Sensitivity Analyses

In the sensitivity analyses where no prior depression was defined by endorsing clinically significant symptoms, we observed a similar pattern of estimated latent classes (Supplemental Table 3). The same pattern of results emerged in the LCMM that further adjusted for neighborhood socioeconomic characteristics (Supplemental Table 4).

### Conclusions

In this longitudinal study conducted on a geographically-diverse cohort of older women with thorough neuropsychological assessments and no prior depression, we used latent-class mixed models to identify six significant clusters of women who exhibited similar trajectories

of depressive symptoms assessed annually up to 11-years. Long-term exposure to ambient PM<sub>2.5</sub> estimated at the residential locations was associated with increased depressive symptoms, especially among the two clusters of older women who experienced sustained-mild and moderate depressive symptoms. This association was robust after adjusting for multiple potential confounders and other risk factors for late-life depression. Our analyses also showed that dementia risk did not fully explain this association. In a subsample of women who underwent structural brain MRI imaging, differences in their volumetric measures of normal-appearing white matter, prefrontal cortex, hippocampus, and small-vessel-ischemic-disease could not explain the observed increase in depressive symptoms associated with long-term PM<sub>2.5</sub> exposure. The neuropathological processes underlying the putative neurotoxic effect of PM<sub>2.5</sub> exposure in late life remains elusive, but are likely independent of cerebrovascular damage.

Our study addresses the limitations of previous research on air pollution and emotional health (see Supplemental Table 5) in several ways. Null associations between long-term PM<sub>2.5</sub> exposure and depressive symptoms of community-dwelling populations were reported in two longitudinal studies (18,19), one conducted in Boston (aged 65 years) and the other across the U.S. (aged 57–85 years). Neither study excluded individuals with a prior history of depression. This restriction not only ensures the study outcome most relevant to late-life depressive symptoms, but also helps reduce the potential confounding by other environmental exposures (including early/mid-life PM<sub>2.5</sub>) and genetic associations presumably affecting emotional health before the baseline. With the depressed symptoms only assessed twice, investigators were unable to characterize the longitudinal trajectories, thus missing the opportunity to examine the potential heterogeneity of exposure effects. Kim et al. (42) reported that long-term PM<sub>2.5</sub> exposure increased the risk of major depression amongst the general population of Seoul, Republic of Korea (15–79 years old). However, participants were followed for two years, and the difference in early-life and late-life depression was disregarded. Only aggregated measures of air pollution exposure were used in their analyses that did not adequately account for confounding by socioeconomic status. Also, their ascertainment for depressive disorders was based on diagnostic codes in medical claims data, an approach known to suffer from outcome misclassification with only modest accuracy (43) and low sensitivity (44). Kioumourtzoglou et al. (17) carefully excluded those with reported prior depression from an US-nationwide cohort of older women (aged 66.6±7.6). Their observed association, however, was sensitive to different outcome definitions and the inclusion of subjects with comorbid conditions presumably associated with late-life depressive symptoms. Although clinical entities of depression often represent arbitrarily defined categories with more extreme manifestations of the continuum of depressive symptoms, the above-mentioned methodological limitations point to the need for future studies with high-quality longitudinal data on outcome (including diagnostic interviews) and comorbidities to examine the association between PM<sub>2.5</sub> and depressive disorders.

Our study adds novel data to the emerging field of environmental neurosciences of air pollution and brain aging. Previous studies have shown long-term PM<sub>2.5</sub> exposure may increase risk of dementia (10), including our report based on WHIMS data (11). Empirical data from WHISCA also showed that GDS-15 trajectories predicted the risk for dementia

(*data not shown*) (45). Taken together, all these data point to two interesting possibilities. First, if late-life depressive symptoms represent a prodromal neuropsychiatric manifestation of dementia (46), findings of the present study suggest that PM<sub>2.5</sub> may perturb the common neuropathological processes leading to the trajectories with sustained depressive symptoms and an increased risk for dementia. Second, it is also possible both depressive symptoms and dementia have shared risk factors, and late-life exposure to PM<sub>2.5</sub> may represent an example of such common environmental causes acting upon overlapping or distinctive pathways to compromise emotion regulation abilities.

No previous epidemiologic data were available to help elucidate the neural basis of increased depressive symptoms associated with air pollution exposure. In the informative subsample of older women participating in both WHISCA and WHIMS-MRI, we found both the global PM<sub>2.5</sub> effect and the class-specific exposure effect (especially among those with sustained-moderate symptoms) remained statistically significant after we further accounted for various MRI-measured regional brain volumes (e.g., prefrontal cortex, normally-appearing white matter) with negative associations with PM<sub>2.5</sub> exposure. These results underscore the need for future studies to investigate other brain regions and neural networks vulnerable to the PM<sub>2.5</sub> neurotoxicity that may predispose older people to increased depressive symptoms in late life.

The findings in this study have potential important policy and clinical implications. These findings emphasize the importance of enforcing environmental regulations (e.g., the Clean Air Act) and initiatives for other public policies to minimize the adverse impact of ambient air pollutants on older adults. Clinicians might consider air pollution as a possible contributing factor to their patients' depressive symptoms and can provide relevant psychoeducation to their patients.

We recognize several limitations of our study. First, although the PM<sub>2.5</sub> spatiotemporal model was statistically cross-validated (average Pearson's  $R^2=0.70$ ) (28), the resulting exposure estimates were still subject to measurement errors. However, such estimation errors are likely non-differential and tend to attenuate the observed associations. Second, we only studied PM<sub>2.5</sub> exposure as regional air pollutants, with no data on its chemical constituencies, emission sources, or possible interactions with other pollutant mixtures. Third, structural MRI data was only available for a subsample of women in our study and was cross-sectional. Fourth, although the application of the latent process transformation to GDS-15 scores was necessary statistically, the clinical interpretation of our findings may not be straightforward. However, given the informative comparison made between the adverse PM<sub>2.5</sub> effect and several clinical comorbidities (Figure 2), we advocate for more use of LCMM with latent process score for studying clinically-relevant questions on late-life depressive symptoms in future research. Fifth, our analyses of brain MRI data were largely based on region-of-interest with only volumetric measures of global regions – an approach that may not have the sufficient sensitivity and specificity to identify areas vulnerable to PM<sub>2.5</sub> neurotoxicity or critical to emotion regulation. We only analyzed data from a single structural MRI and thus limits our ability to make inferences about the extent to which PM<sub>2.5</sub> exposure is associated with change in structural MRI variables and depressive symptoms over time. Future research needs to utilize repeated MRI measures to examine

whether there are PM<sub>2.5</sub> associated changes in brain volumes based on depressive symptom trajectories. Lastly, future studies should consider high-dimensional analyses, including the use of machine learning (47) that may offer a better approach to understanding the likely complex brain structure and neural networks linking PM<sub>2.5</sub> exposure with depressive symptoms in late life.

This study had several major strengths. First, women were followed over a long period up to 11 years with a maximum of nine assessments of their depressive symptoms. Second, we employed the sophisticated LCMM modeling approach, which allowed us to identify the trajectories of depressive symptoms as the basis for examining the heterogeneity of PM<sub>2.5</sub> exposure effects. Third, we had access to well-validated data on dementia and MRI-based measures of regional brain volumes and cerebrovascular damage, which enabled the first step to explore the potential contributors to the neuropathological processes linking late-life exposure to PM<sub>2.5</sub> with GDS-15 trajectories. Lastly, the rich, comprehensive covariate data from the WHIMS cohort offered a unique population context for optimal study design that excludes individuals with prior depression and rigorously assess and adjust for potential confounding.

In conclusion, among older women with no prior depression or dementia, long-term exposure to ambient fine particles is a novel environmental risk factor for sustained depressive symptoms, independent of other known risk factors for late-life depression. Future studies need to better understand the biological underpinnings of this association.

## Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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**Highlights****What is the primary question addressed by this study?**

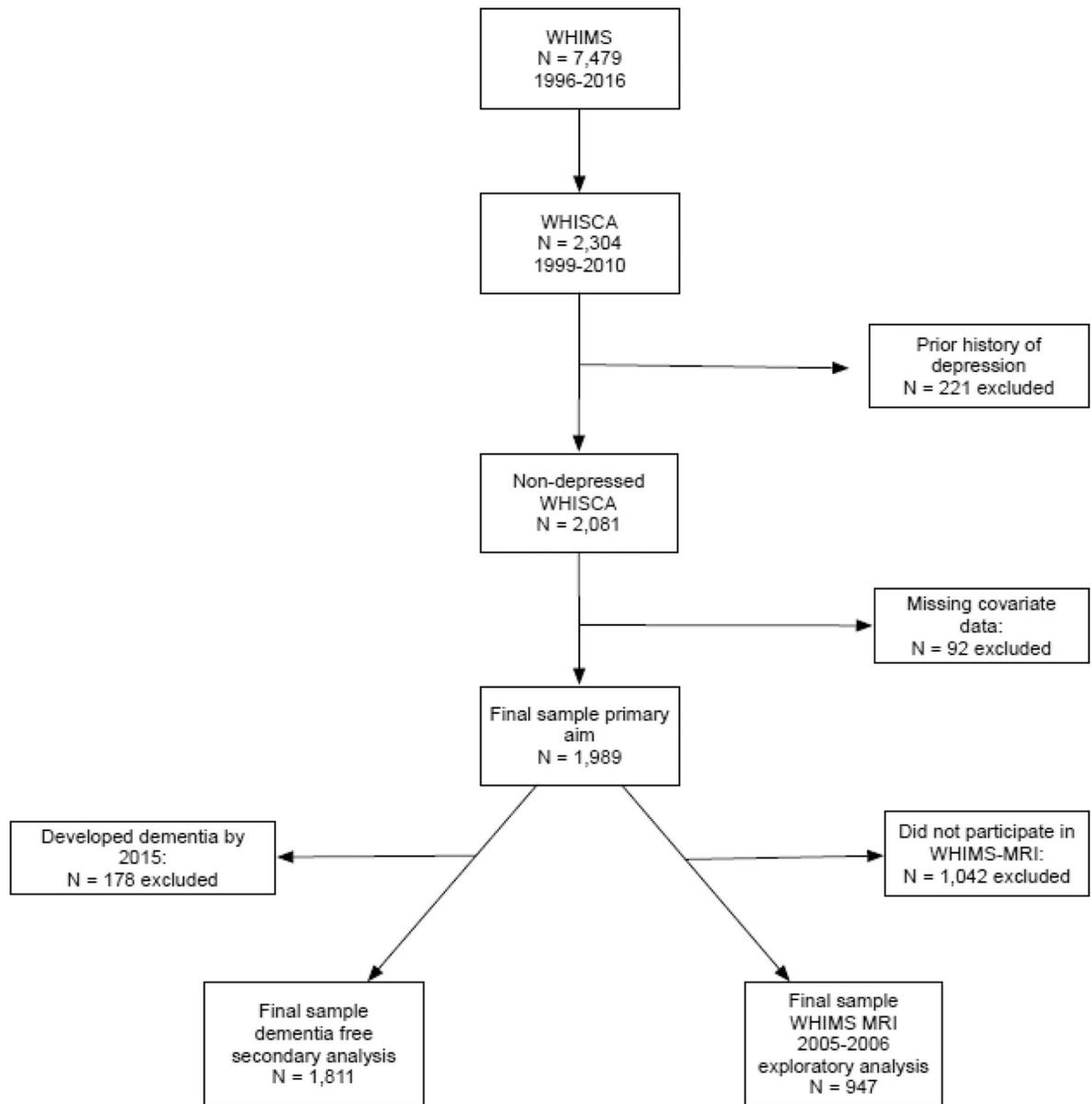
Is late-life exposure to PM<sub>2.5</sub> (particulate matter with aerodynamic diameters < 2.5 μm) associated with depressive symptoms in older women without prior depression or dementia and if so, how?

**What is the main finding of this study?**

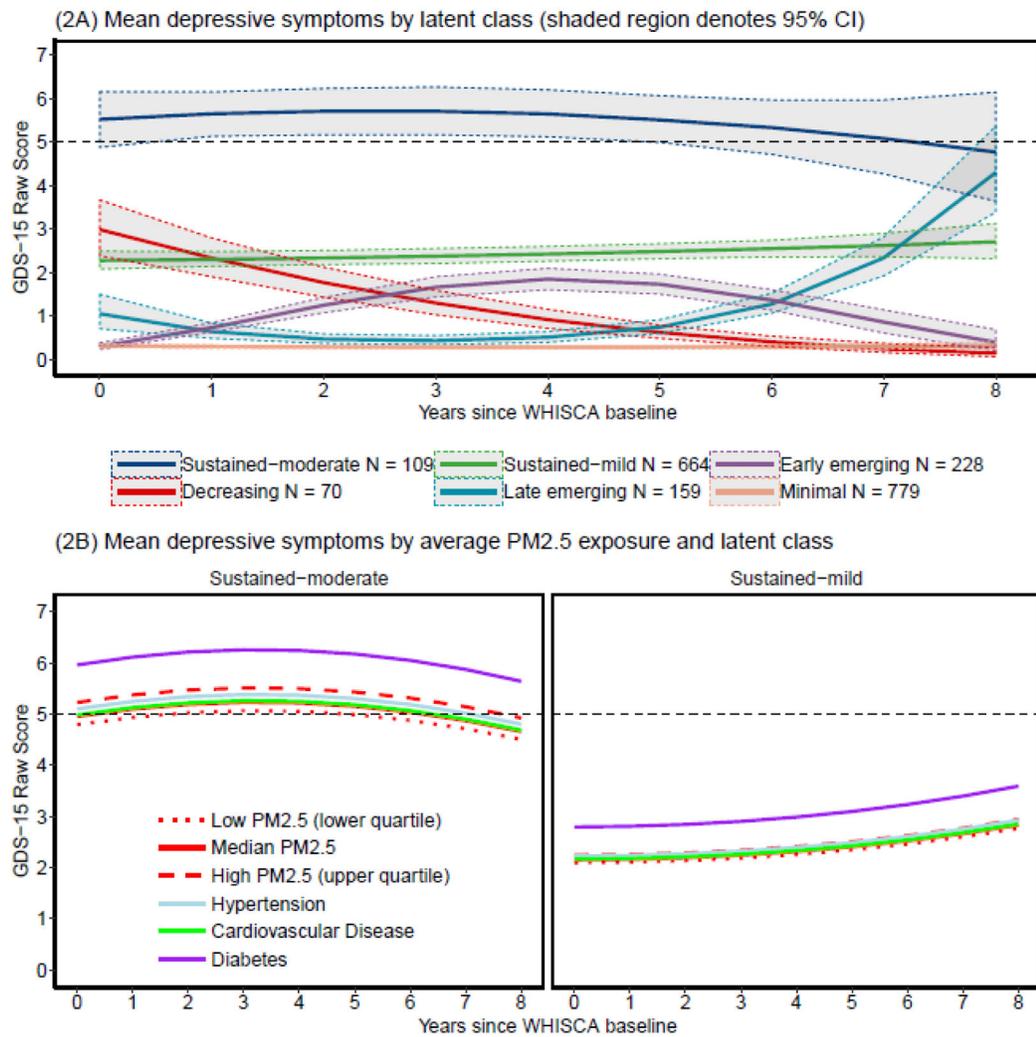
In this 11-year longitudinal study (n= 1,989), residing in locations with higher ambient PM<sub>2.5</sub> increased the annually-assessed depressive symptoms, especially among women with sustained mild or moderate symptoms. Dementia risk only partially explained this association, which was independent of MRI-measured brain volumes and small-vessel-ischemic-disease.

**What is the meaning of this study?**

In older women, late-life PM<sub>2.5</sub> exposure is a novel environmental risk factor of depressive symptoms, but the neurobiological mechanisms are unclear.



**Figure 1.**  
Flowchart of study participants included in the analyses



**Figure 2.** Graph of the estimated mean score of the 15-item Geriatric Depression Scale with 95% confidence intervals over time by latent class (panel A). The trajectories of depressive symptoms affected by exposure were depicted over time by high (upper quartile), mean, and low (lower quartile) PM<sub>2.5</sub> exposure for women with sustained depressive symptoms (panel B). The reference trajectories of estimated average scores on 15-item Geriatric Depression Scale for women with hypertension, cardiovascular disease, or diabetes were plotted as reference groups with assumed exposures to median PM<sub>2.5</sub> levels. A dotted reference line was added at the GDS-15 score of 5 in order to provide a benchmark for clinically significant depressive symptoms.

**Table 1.**Comparison of Estimated PM<sub>2.5</sub> Exposure by Baseline Population Characteristics (N=1,989).

Population Characteristics	N	Distribution of time-varying 3-year average PM <sub>2.5</sub> <sup>a</sup>		p <sup>b</sup>
		Mean ± SD (25 <sup>th</sup> percentile, median, 75 <sup>th</sup> percentile)		
<b>Overall</b>	1,989	12.61 ± 2.58 (10.74, 12.03, 14.19)		
<b>Region</b>				<0.01
Northeast	418	12.18 ± 1.33 (11.1, 12.1, 13.2)		
South	299	12.21 ± 1.77 (10.6, 11.8, 14)		
Midwest	772	12.85 ± 2.43 (10.7, 12.3, 15.4)		
West	500	12.84 ± 3.69 (10.7, 11.8, 14.9)		
<b>Ethnicity</b>				<0.01
African-American	115	15.23 ± 2.39 (13.8, 14.7, 16.6)		
Hispanic White	23	12.74 ± 2.45 (11.1, 11.9, 14.1)		
White (not Hispanic)	1802	12.41 ± 2.47 (10.7, 11.8, 13.9)		
Other or Missing	49	13.81 ± 3.41 (11.4, 12.9, 16.2)		
<b>Education</b>				0.02
Less than high school	89	12.48 ± 2.55 (10.9, 11.7, 14.1)		
High school	431	12.32 ± 2.24 (10.6, 11.7, 13.6)		
More than high school	1469	12.7 ± 2.67 (10.8, 12.1, 14.3)		
<b>Employment</b>				0.43
Currently working	383	12.49 ± 2.61 (10.7, 12, 14)		
Not working	190	12.79 ± 2.62 (10.8, 12.2, 14.7)		
Retired	1416	12.62 ± 2.56 (10.8, 12, 14.2)		
<b>Income (USD)</b>				<0.01
< 9,999	422	12.34 ± 2.57 (10.6, 11.5, 14.1)		
10,000–34,999	625	12.58 ± 2.51 (10.7, 12.1, 14.1)		
35,000–49,999	430	12.6 ± 2.53 (10.7, 12, 14)		
50,000–74,999	277	12.82 ± 2.71 (11, 12.2, 14.2)		
75,000 or more	176	13.19 ± 2.75 (11, 12.5, 14.9)		
Don't know	59	12.27 ± 2.3 (10.7, 11.9, 13.5)		
<b>Lifestyle</b>				
<b>Smoking status</b>				0.44
Never smoked	1106	12.64 ± 2.61 (10.7, 12, 14.3)		
Past smoker	775	12.61 ± 2.51 (10.8, 12.1, 14)		
Current Smoker	108	12.31 ± 2.72 (10.5, 11.7, 14.1)		
<b>Alcohol use</b>				0.07
Non-drinker	246	12.99 ± 2.7 (10.7, 12.5, 14.8)		
Past drinker	347	12.57 ± 2.54 (10.7, 12, 14.4)		
Less than 1 drink per day	1148	12.52 ± 2.55 (10.7, 12, 13.9)		
More than 1 drink per day	248	12.69 ± 2.61 (10.9, 12, 14.2)		

Distribution of time-varying 3-year average PM <sub>2.5</sub> <sup>a</sup>				
Population Characteristics	N	Mean ± SD		p <sup>b</sup>
		(25 <sup>th</sup> percentile, median, 75 <sup>th</sup> percentile)		
<b>Moderate or strenuous physical activities 20 minutes</b>				0.42
No activity	1128	12.65 ± 2.64 (10.7, 12.1, 14.3)		
Some activity	107	12.73 ± 2.35 (10.8, 12.2, 14.6)		
2–4 episodes/wk	413	12.64 ± 2.48 (10.8, 12, 14.2)		
4 episodes/week	341	12.4 ± 2.56 (10.6, 11.8, 13.7)		
<b>Physical Health</b>				
<b>Hypertension</b>				0.24
No	1261	12.56 ± 2.57 (10.7, 12, 14.2)		
Yes	728	12.7 ± 2.6 (10.8, 12.1, 14.3)		
<b>High Cholesterol</b>				0.14
No	1641	12.57 ± 2.57 (10.7, 12, 14.2)		
Yes	348	12.8 ± 2.6 (10.8, 12.2, 14.3)		
<b>Diabetes</b>				0.84
No	1890	12.61 ± 2.59 (10.7, 12, 14.2)		
Yes	99	12.56 ± 2.34 (10.9, 12.1, 13.8)		
<b>Cardiovascular disease</b>				0.03
No	1680	12.56 ± 2.56 (10.7, 12, 14.1)		
Yes	309	12.91 ± 2.65 (10.9, 12.4, 14.8)		
<b>Hormone treatment ever</b>				0.86
No	1083	12.6 ± 2.49 (10.7, 12.1, 14.1)		
Yes	906	12.62 ± 2.68 (10.8, 12, 14.3)		

<sup>a</sup>PM<sub>2.5</sub> represents the distribution of the individual-level summary of all time-varying 3-year exposures aggregated from the daily exposure levels estimated at each residential location using the spatiotemporal model.

<sup>b</sup>p values estimated from ANOVA F-tests or t-tests

**Table 2.**

Comparison of Baseline Population Characteristics by Latent Class<sup>a</sup> of Depressive Symptoms in the Women's Health Initiative Study on Cognitive Aging (N = 1,989).

Population Characteristics	Resilient	Early emerging	Late emerging	Sustained-Mild	Decreasing	Sustained-Moderate	p <sup>c</sup>
	N = 779	N = 228	N = 159	N = 644	N = 70	N = 109	
	Mean (sd)						
	or	or	or	or	or	or	
	% <sup>b</sup> (N)						
<b>Age (years)</b>	73.00 (3.56)	73.37 (3.82)	72.72 (3.45)	73.83 (4.02)	72.89 (3.61)	73.87 (3.71)	<.01
<b>GDS-15<sup>d</sup></b>	.25 (.58)	.09 (.3)	1.36 (1.43)	2.12 (1.57)	3.09 (2.25)	5.79 (2.5)	<.01
<b>Region</b>							.08
Northeast	39% (163)	10% (43)	8% (33)	34% (144)	3% (12)	6% (23)	
South	37% (110)	12% (37)	10% (31)	33% (100)	3% (9)	4% (12)	
Midwest	40% (310)	14% (109)	8% (59)	29% (220)	4% (30)	6% (44)	
West	39% (196)	8% (39)	7% (36)	36% (180)	4% (19)	6% (30)	
<b>Ethnicity</b>							<.01
African-American	34% (39)	10% (11)	6% (7)	36% (41)	3% (3)	12% (14)	
Hispanic White	17% (4)	17% (4)	4% (1)	35% (8)	4% (1)	22% (5)	
White (not Hispanic)	40% (713)	12% (212)	8% (150)	32% (577)	4% (64)	4% (86)	
Other or Missing	47% (23)	2% (1)	1% (1)	37% (18)	4% (2)	8% (4)	
<b>Education</b>							<.01
Less than high school	24% (21)	16% (14)	1% (1)	45% (40)	3% (3)	11% (10)	
High school	38% (163)	11% (46)	9% (38)	32% (137)	4% (16)	7% (31)	
More than high school	41% (595)	11% (168)	8% (120)	32% (467)	3% (51)	5% (68)	
<b>Employment</b>							.01
Currently working	39% (148)	10% (37)	12% (47)	32% (122)	3% (11)	5% (18)	
Not working	40% (76)	10% (18)	2% (4)	40% (76)	4% (7)	5% (9)	
Retired	39% (555)	12% (173)	7% (108)	32% (446)	4% (52)	6% (82)	
<b>Annual income (dollars)</b>							<.01
< 9,999	31% (129)	12% (50)	8% (32)	39% (165)	2% (9)	9% (37)	
10,000–34,999	38% (235)	10% (65)	8% (52)	36% (224)	3% (18)	5% (31)	
35,000–49,999	43% (185)	13% (58)	8% (34)	26% (110)	4% (18)	6% (25)	
50,000–74,999	48% (132)	9% (26)	10% (28)	24% (66)	5% (13)	4% (12)	
75,000 or more	45% (79)	14% (25)	5% (9)	31% (55)	4% (7)	1% (1)	
Don't know	32% (19)	7% (4)	7% (4)	41% (24)	8% (5)	5% (3)	
<b>Lifestyle</b>							
<b>Smoking status</b>							0.55
Never smoked	40% (447)	12% (130)	8% (89)	30% (337)	3% (38)	6% (65)	
Past smoker	38% (292)	12% (91)	8% (62)	34% (266)	3% (28)	5% (36)	

	<b>Resilient</b>	<b>Early emerging</b>	<b>Late emerging</b>	<b>Sustained-Mild</b>	<b>Decreasing</b>	<b>Sustained-Moderate</b>	
	<b>N = 779</b>	<b>N = 228</b>	<b>N = 159</b>	<b>N = 644</b>	<b>N = 70</b>	<b>N = 109</b>	
	<b>Mean (sd)</b>						
	<b>or</b>	<b>or</b>	<b>or</b>	<b>or</b>	<b>or</b>	<b>or</b>	
<b>Population Characteristics</b>	<b>%<sup>b</sup> (N)</b>	<b>p<sup>c</sup></b>					
Current Smoker	37% (40)	6% (7)	7% (8)	38% (41)	4% (4)	7% (8)	
<b>Alcohol use</b>							0.03
Non-drinker	42% (103)	14% (34)	8% (20)	29% (71)	4% (11)	3% (7)	
Past drinker	32% (112)	9% (32)	10% (33)	39% (134)	3% (10)	7% (26)	
Less than 1 drink / day	39% (449)	12% (137)	8% (89)	32% (364)	4% (42)	6% (67)	
More than 1 drink / day	46% (115)	10% (25)	7% (17)	30% (75)	3% (7)	4% (9)	
<b>Moderate or strenuous physical activities 20 minutes</b>							<.01
No activity	35% (394)	10% (113)	8% (93)	37% (415)	3% (35)	7% (78)	
Some activity	37% (40)	17% (18)	12% (13)	25% (27)	3% (3)	6% (6)	
2–4 episodes/wk	44% (180)	13% (53)	4% (18)	30% (125)	5% (19)	4% (18)	
episodes/week	48% (165)	13% (44)	10% (35)	23% (77)	4% (13)	2% (7)	
<b>Physical Health</b>							
<b>Hypertension</b>							<.01
No	42% (532)	11% (143)	8% (98)	29% (368)	4% (52)	5% (68)	
Yes	34% (247)	12% (85)	8% (61)	38% (276)	2% (18)	6% (41)	
<b>High Cholesterol</b>							.53
No	40% (657)	11% (187)	8% (130)	32% (526)	3% (54)	5% (87)	
Yes	35% (122)	12% (41)	8% (29)	34% (118)	5% (16)	6% (22)	
<b>Diabetes</b>							<.01
No	40% (753)	12% (224)	8% (151)	32% (599)	4% (69)	5% (94)	
Yes	26% (26)	4% (4)	8% (8)	45% (45)	1% (1)	15% (15)	
<b>Cardiovascular disease</b>							.01
No	41% (685)	11% (187)	8% (133)	31% (527)	4% (62)	5% (86)	
Yes	30% (94)	13% (41)	8% (26)	38% (117)	3% (8)	7% (23)	
<b>Hormone treatment ever</b>							.30
No	41% (447)	11% (123)	8% (89)	30% (329)	4% (38)	5% (57)	
Yes	37% (332)	12% (105)	8% (70)	35% (315)	4% (32)	6% (52)	
<b>Baseline exposure</b>							
PM <sub>2.5</sub> (3-year average)	12.78 (2.4)	12.83 (2.4)	12.64 (2.16)	12.66 (2.46)	12.56 (2.38)	13.14 (2.52)	.42

<sup>a</sup> Group membership estimated from the latent class model that includes the initial level of depressive symptoms with both fixed effects and random intercept, linear and quadratic change of follow-up time, initial WHISCA age, and attrition (yes/no if participant dropped out at the subsequent assessment time) to account for heterogeneity in the longitudinal profiles of depressive symptoms.

<sup>b</sup> Percentage reported above represent row percentages.

<sup>c</sup> p values calculated from chi-square or fisher exact tests for categorical variables, from ANOVA F-tests for continuous variables

<sup>d</sup> GDS-15 = Geriatric Depression Scale-15

**Table 3.**

Effect of Time-Varying 3-Year Average PM<sub>2.5</sub> Exposure on Depressive Symptoms Measured by the 15-Item Geriatric Depression Scale in the Women's Health Initiative Study of Cognitive Aging<sup>a</sup>

	PM <sub>2.5</sub> Effect Estimate						
	Full sample N = 1,989			Dementia-free sample N=1,811			
	$\beta^b$	SE <sup>c</sup>	p	$\beta^b$	SE <sup>c</sup>	p	
<b>Model 1: Global main effect</b>							
Exposure effect estimate common to all classes	<b>.043</b>	<b>.021</b>	<b>.039</b>	Exposure effect estimate common to all classes	<b>.038</b>	<b>.014</b>	<b>.006</b>
<b>Model 2: Class-specific exposure effect</b>							
Minimal (n=802)	.021	.026	.426	Minimal (n=737)	.014	.020	.489
Early-emerging (n=214)	-.005	.124	.968	Early-emerging (n=199)	-.008	.042	.855
Late-emerging (n=158)	.168	.106	.113	Late-emerging (n=141)	<b>.175</b>	<b>.050</b>	<b>&lt;.001</b>
Sustained-mild (n=626)	<b>.096</b>	<b>.042</b>	<b>.023</b>	Sustained-mild (n=568)	<b>.086</b>	<b>.026</b>	<b>&lt;.001</b>
Decreasing (n=67)	-.141	.104	.176	Decreasing (n=62)	-.078	.061	.201
Sustained-moderate (n=122)	<b>.171</b>	<b>.076</b>	<b>.024</b>	Sustained-moderate (n=104)	<b>.151</b>	<b>.058</b>	<b>.009</b>

<sup>a</sup>All models adjust for age at WHISCA baseline, race/ethnicity, attrition, geographic region of residence, education, household income, employment status, baseline global cognitive ability, lifestyle factors (smoking, alcohol use, physical activities) and clinical characteristics (use of hormone treatment; hypercholesterolemia, hypertension, diabetes, and history of cardiovascular disease).

<sup>b</sup> $\beta$  = the average increase in transformed 15-item Geriatric Depression Scale score per increase of one interquartile range (3.45  $\mu\text{g}/\text{m}^3$ ) of time-varying 3-year average exposure to PM<sub>2.5</sub>

<sup>c</sup>SE = the standard error of the estimated effect of PM<sub>2.5</sub> on 15-item Geriatric Depression Scale.

Bolded estimates denote  $p < .05$